Bankruptcy and small firms' access to credit

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We investigate how personal bankruptcy law affects small firms' access to credit. When a firm is unincorporated, its debts are personal liabilities of the firm's owner, so that lending to the firm is legally equivalent to lending to its owner. If the firm fails, the owner has an incentive to file for personal bankruptcy, since the firm's debts will be discharged. The higher the exemption level, the greater the incentive to file for bankruptcy. We test the model and find that if small firms are located in states with unlimited rather than low homestead exemptions, they are more likely to be denied credit, and when loans are made, they are smaller and interest rates are higher. Results for noncorporate versus corporate firms suggest that lenders often disregard small firms' organizational status in making loan decisions.

1. Introduction

■ Small businesses are the primary source of new jobs in the U.S. economy. From 1990 to 1995, businesses with fewer than 500 employees accounted for 76.5% of net new jobs. But small businesses have a very high turnover rate compared to large businesses. Over 13% of U.S. jobs in 1995 were in firms that did not exist before 1990, and over 12% of jobs in 1990 were in firms that had ceased to exist by 1995.¹ Despite the importance and the complexity of small business as a contributor to the U.S. economy, there has been surprisingly little academic research on the economic environment faced by small business owners or the effects of policy variables on small business success.

In this article we investigate how personal bankruptcy law affects small firms' access to credit. It is well known that the United States has separate bankruptcy procedures for individuals versus corporations. What is less well known is that personal bankruptcy procedures also apply to

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¹ See U.S. Small Business Administration (1998).

small firms. When a firm is noncorporate, its debts are personal liabilities of the firm's owner, so that lending to the firm is legally equivalent to lending to its owner. If the firm fails, the owner can file for bankruptcy and her business and unsecured personal debts will be discharged. When a firm is a corporation, limited liability implies that the owner is not legally responsible for the firm's debts. However, lenders to small corporations often require that the owner guarantee the loan and may also require that the owner give the lender a second mortgage on her house. This wipes out the owner's limited liability for purposes of the particular loan and makes small corporate firms into corporate/noncorporate hybrids. Thus, personal bankruptcy law applies to noncorporate firms and may also apply to small corporate firms.²

In this article we test whether variation in personal bankruptcy exemptions across U.S. states affects small firms' access to credit, using the 1993 *National Survey of Small Business Finance* (NSSBF). We find that small businesses are more likely to be denied credit if they are located in states with high rather than low homestead exemptions and that, if they receive loans, the loans are smaller and interest rates are higher. We also find that bankruptcy exemption levels affect both noncorporate and corporate firms, suggesting that lenders often ignore a small business's organizational status when making their loan decisions.

The remainder of the article is organized as follows. After a brief literature review, Section 3 discusses personal bankruptcy law. Section 4 discusses a simple model of small business credit markets. Section 5 presents the empirical results, and Section 6 concludes.

2. Literature review

■ We know of only one article that examines the effect of personal bankruptcy law on business credit markets. Scott and Smith (1986) examined the effect of the new U.S. Bankruptcy Code, adopted in 1978, on business credit markets. They argued that adoption of the code caused the cost of business loans to increase and that lenders raised interest rates in response. They found support for this hypothesis in their empirical work. However, their study examined only the net effect on interest rates of many changes adopted simultaneously as part of the 1978 code, all of which applied uniformly over the United States. Our study, in contrast, uses cross-section variation in bankruptcy exemption levels across U.S. states to examine whether exemption levels affect the availability of small business credit.

On the personal bankruptcy side, Gropp, Scholz, and White (1997) investigated how varying bankruptcy exemption levels across states affect markets for nonbusiness loans. They found that in states with higher exemption levels, applicants were more likely to be turned down for credit, but demand for loans increased. Overall they found that higher bankruptcy exemption levels shift credit from households with low assets to those with high assets, since lenders are willing to accommodate the increased demand of the latter group but not the former. Berkowitz and Hynes (1999) and Lin and White (2001) reexamined this issue for mortgage loans.

Petersen and Rajan (1994, 1995) examine small business credit markets using earlier versions of the NSSBF. They are mainly concerned with examining the effects of long-term relationships between firms and banks and the effects of concentration in local banking markets on interest rates and availability of business credit. More recent research on banking relationships, using a later version of the NSSBF, includes Cole (1998) and Cole, Goldberg, and White (2000). Cavalluzzo, Cavalluzzo, and Wolken (2002) examine patterns of race discrimination in banking relationships. There are also a number of theoretical models that examine the effects of bankruptcy on credit markets.³

² Sullivan, Warren, and Westbrook (1989) surveyed a sample of individuals who filed for bankruptcy during the 1980s and estimated that around 20% had debts from a failed business. This is nearly double the proportion of all U.S. households that have self-employment income.

³ Bester (1994) predicts that firms are less likely to default when loans are collateralized, and Hart and Moore (1998) predict that firms are more likely to default when they have a single creditor rather than multiple creditors and when the liquidation value of the firm's assets is lower. Unfortunately, our data do not allow these predictions to be tested, because default can occur anytime during the seven years prior to the survey, while other firm characteristics are for the time of the survey only.

3. Bankruptcy law and small firms

When unincorporated firms fail, their owners typically have high debt levels, much of which consists of debts of the failed firm. Owners have an incentive to file for bankruptcy, both because their unsecured personal and business debts will be discharged and because creditors must immediately terminate collection efforts and legal actions to obtain repayment. Under the Chapter 7 bankruptcy procedure, debtors' future earnings are completely exempt from the obligation to repay prebankruptcy debt, but they must turn over any assets they own above an exemption level to the bankruptcy trustee, who uses these assets to repay debt. When debtors file under Chapter 7, they cannot file again for six years.

While bankruptcy is a matter of federal law and the procedure is uniform across the country, Congress gave the states the right to set their own bankruptcy exemption levels, and these vary widely. Most states have several types of exemptions: for equity in owner-occupied principal residences (the homestead exemption), for equity in cars, for cash, and for various types of goods (furniture, clothing, cooking utensils, farm implements, family bibles, tools of the trade, sometimes a horse, etc.). In most states, the homestead exemption is the largest, and other exemptions are small. Debtors therefore have an incentive to convert cash or financial assets into additional home equity by repaying part of their mortgage loans before filing for bankruptcy, if their home equity is less than the homestead exemption.⁵

There is also a second personal bankruptcy procedure, Chapter 13, and debtors are allowed to choose between them. Under Chapter 13, debtors must propose a plan to use some of their future earnings to repay part or all of their debt, but all of their assets are exempt. This procedure is generally less favorable to business owners than Chapter 7, because failed entrepreneurs often have no nonexempt assets and because having an obligation to repay past debt from future earnings would make it difficult to start a new business. Because creditors are entitled to receive no less in Chapter 13 than they would have received in Chapter 7 (13 U.S.C. Section 1325(a)(4)), exemption levels are likely to have similar effects on credit markets regardless of the chapter that business owners would choose if they filed for bankruptcy.⁶

Now consider the possibility that the firm is incorporated. Corporate firms are legally separate from their owners, so owners are not personally responsible for the debts of their corporations. Holding everything else constant, this means that small corporations are less creditworthy than small noncorporate firms, because the former have only the corporation's assets to back up business debt, while the latter have both the firm's assets and the owner's personal assets. Lenders also know that owners of small corporations can easily shift assets between their personal accounts and their corporations' accounts, so that lenders may not view the corporate/noncorporate distinction as meaningful for small firms. In making loans to small corporations, lenders therefore may require that owners personally guarantee the loans. This abolishes the legal distinction between corporations and their owners for purposes of the particular loan and puts the owner's personal assets at risk to repay the loan.

4. A stylized model

Suppose an entrepreneur owns a noncorporate firm and wishes to invest in a risky project. In period 1, the entrepreneur has wealth W. She can apply for a loan, D, to be used for a project with risky return, R. Assuming that the loan is made, it will have an interest rate of r and be due

⁴ One might question whether potential entrepreneurs would be familiar with bankruptcy law and bankruptcy exemptions. We found that self-help manuals such as *Legal Guide to Starting and Running a Small Business*, *Vol. 1*, contain a clear explanation of bankruptcy. See Steingold (1999).

⁵ White (1998b) discusses various strategies for sheltering assets in bankruptcy.

⁶ Creditors do not receive more in Chapter 13 than Chapter 7 because debtors choose to file under whichever Chaper minimizes their liability. An additional restriction that makes Chapter 13 less attractive than Chapter 7 to small business owners is that unsecured debt discharged in Chapter 13 cannot be higher than \$250,000. There is no limit in Chapter 7.

in period 2. The owner's wealth in period 2 if she repays the loan will be W + R - D(1+r), which we denote Z. We further assume that Z has strictly positive density, f(Z).

In period 2, the owner may file for bankruptcy under Chapter 7. Denote the combined value of the various bankruptcy exemptions in the owner's state of residence as X. Also suppose the cost of filing for bankruptcy is C, where C < D(1+r). C is assumed to include both the out-of-pocket cost of filing and the cost of reduced access to credit in the future. If the owner files for bankruptcy, then the debt of D(1+r) will be discharged, but she must use any wealth that exceeds the exemption level, or $\max[Z-X-C,0]$, to repay the debt. Let Z^* be the level of period 2 wealth at which owners are indifferent between filing and not filing for bankruptcy, so that they file for bankruptcy if $Z < Z^*$ and repay in full otherwise. In the region around Z^* , owners' net wealth is Z-D(1+r) if they do not file for bankruptcy and X if they do (assuming that they pay the cost of bankruptcy C before filing). Therefore they are indifferent between filing and not filing at the wealth level $Z^* = X + D(1+r)$. This means that owners' net wealth in period 2 has three regions: a high region in which owners repay the debt in full and their net wealth is Z-D(1+r); a middle region in which owners file for bankruptcy, repay Z-X-C in bankruptcy, and have net wealth of X; and a low region in which owners file for bankruptcy, repay nothing, and have net wealth Z-C. The dividing point between the middle and low regions is Z=X+C.

A representative lender maximizes expected profits,

$$\int_{X+C}^{X+D(1+r)} (Z-X-C)f(Z)dZ + \int_{X+D(1+r)}^{\infty} D(1+r)f(Z)dZ, \tag{1}$$

with respect to the interest rate they are willing to offer. The first term on the left-hand side is the expected value of repayment if owners file for bankruptcy, and the second term is the expected value of repayment if owners avoid bankruptcy.

Lenders realize that the amount of debt that owners apply for depends on the interest rate and the exemption level, so that D = D(X, r). Solving for the first-order condition of (1) with respect to r and assuming for simplicity that C = 0, we get

$$\int_{X+D(X,r)(1+r)}^{\infty} [D_r'(X,r)(1+r) + D(X,r)]f(Z)dZ = 0.$$
 (2)

Since f(Z) is positive, this becomes $D(X,r) = -D'_r(X,r)(1+r)$. To ensure that both D and (1+r) are nonnegative, the sign of D'_r must be negative, i.e., entrepreneurs' demand for debt must be negatively related to the interest rate.

As an example, suppose the demand curve for debt is additive, $D(X,r) = \alpha(X) - \beta r$, where β is a positive constant. Solving for the optimal interest rate, $r = (1/2)(\alpha(X)/\beta - 1)$. The slope of the relationship between the interest rate and exemption level is determined by the sign of $\alpha(X)$. Moreover, substituting r into the demand curve, we find $D = (1/2)(\beta + \alpha(X))$. Thus, loan size should vary in the same way as the interest rate in response to changes in the exemption.

We expect the sign of $\alpha(X)$ to be generally positive. Bankruptcy provides partial wealth insurance to business owners by transferring some wealth from nonbankruptcy to bankruptcy states. When the exemption level rises, the level of partial wealth insurance increases. Assuming that owners are risk averse, they benefit from the additional wealth insurance, and therefore their demand for loans increases.⁸

The market-clearing condition for this model is that expected return (1) equal lenders' opportunity cost of funds, or

$$D(1+r_f) = D(1+r) \int_{X+D(1+r)}^{\infty} f dZ + \int_{X}^{X+D(1+r)} (Z-X) f dZ,$$
 (3)

⁷ The model would be unchanged if part of the uncertainty in the entrepreneur's period-2 wealth came from other sources, such as risky projects that the entrepreneur invested in before period 1.

⁸ See Fan and White (2003) for a discussion of how bankruptcy provides partial wealth insurance for entrepreneurs and a derivation of the optimal bankruptcy exemption level.

where r_f is the available risk-free rate. Equation (3) implies that the market-clearing interest rate r is positively related to the probability of owners filing for bankruptcy and negatively related to the lender's expected percentage return conditional on bankruptcy. As the exemption level rises, the probability of bankruptcy rises and the expected percentage return in bankruptcy falls (since entrepreneurs repay Z - X if $X < Z < Z^*$), so that the market-clearing interest rate must rise. If the exemption level is high enough, all lending ceases because no interest rate is high enough to clear the market and/or owners' demand for credit falls to zero. Note that in this simple model, owners always file for bankruptcy if the exemption level is unlimited. Therefore, lenders would never lend in states with unlimited exemptions.

In this model, loan market either clears or shuts down completely. But a mixture of both outcomes would occur if the model were extended to allow individual owners' creditworthiness to vary according to some observable characteristic, such as owners' wealth or their firms' wealth in period 1. Then as the exemption level rises, lenders would first reach the no-lending threshold for owners who have the lowest wealth. They would cease lending to low-wealth owners, but continue lending to owners whose wealth is higher. If the exemption level continued to rise, lenders would cease lending to owners having medium wealth and, eventually, to owners having high wealth. Thus as the exemption level rises, lenders gradually cease lending to a higher fraction of potential borrowers (see Longhofer, 1997). If the exemption level were unlimited, then in this simple model lenders would not lend at all, since even high-wealth borrowers would always file for bankruptcy.⁹

Finally, note that although we did not differentiate between the homestead and other exemptions, it is straightforward to show that the predictions are qualitatively the same for both. Berkowitz and Hynes (1999) and Lin and White (2001) present models that differentiate between the two types of exemptions and also differentiate between secured (mortgage) versus unsecured (business) loans.

5. Empirical tests

Our primary data source is the 1993 NSSBF. This survey covers a representative sample of U.S. nonfinancial, nonfarm, for-profit businesses that have fewer than 500 employees. There are approximately 1,750 noncorporate firms and 2,800 corporate firms in the sample.¹⁰

The NSSBF asks managers whether their firms applied for credit within the past three years and, if so, whether the most recent application for credit was denied. Managers are also asked whether they were discouraged from applying for loans at any time during the past three years because they anticipated being turned down. We define a dummy variable, "discouraged/denied," which equals one if managers either were discouraged from applying for credit or applied but were turned down on their most recent application within the past three years. We refer to firms that are discouraged/denied as credit rationed, and we run separate regressions explaining whether corporate and noncorporate firms are credit rationed.¹¹

The main explanatory variables of interest are the homestead and personal property exemptions in the state where the firm is located. The personal property exemption is defined as the sum of the state's exemptions for cash and for equity in vehicles, plus the value of the wildcard exemption. Table 1 shows the two exemptions by state in 1993. The homestead exemption varies widely across states, from zero in one state to unlimited in seven states, with a median value of \$15,000. In contrast, the personal property exemption is both smaller and less variable. The

⁹ Credit rationing is usually associated with models that assume heterogeneous borrowers and asymmetric information concerning borrowers' types. But as the discussion here shows, a form of credit rationing may also occur when all information is common knowledge. See Stiglitz and Weiss (1981) for the original model in the credit-rationing literature.

¹⁰ The NSSBF is produced by the Federal Reserve Board and the Small Business Administration. See Cole and Wolken (1995) for discussion. We use the internal version of the dataset, which identifies the state in which the firm is located.

¹¹ Firms may be both discouraged and denied, because they may have been turned down on their most recent loan application and also discouraged from applying for a loan at some time during the past three years. See below for tests of whether the two types of firms differ.

TABLE 1 1993 Bankruptcy Exemptions by State

	Homestead Exemption (\$)	Personal Property Exemption (\$)
Alabama	10,000	6,000
Alaska	54,000	6,000
Arizona	100,000	3,300
Arkansas	Unlimited	2,900
California	75,000	5,000
Colorado	60,000	2,000
Connecticut	15,000	10,700
D.C.	15,000	10,700
Delaware	15,000	10,700
Florida	Unlimited	2,000
Georgia	10,000	2,800
Hawaii	40,000	42,000
Iowa	Unlimited	10,200
Idaho	100,000	3,000
Illinois	15,000	6,400
Indiana	15,000	8,200
Kansas	Unlimited	40,000
Kentucky	10,000	7,000
Louisiana	15,000	25,000
Massachusetts	100,000	3,350
Maryland	0	11,000
Maine	15,000	3,200
Michigan	15,000	10,700
Minnesota	Unlimited	6,000
Missouri	8,000	3,500
Mississippi	150,000	20,000
Montana	80,000	1,200
North Carolina	20,000	10,000
North Dakota	160,000	12,400
Nebraska	20,000	5,000
New Hampshire	60,000	2,000
New Jersey	15,000	10,700
New Mexico	40,000	9,000
Nevada	95,000	3,000
New York	20,000	9,800
Ohio	10,000	3,600
Oklahoma	Unlimited	6,000
Oregon	20,000	16,200
Pennsylvania	15,000	10,700
Rhode Island	15,000	10,700
South Carolina	15,000	10,700
South Dakota	60,000	4,000
Tennessee	7,500	8,000
Texas	Unlimited	60,000
Utah	10,000	3,000
Virginia	10,000	10,000
Vermont	60,000	21,200
Washington	60,000	7,000
Wisconsin	40,000	4,400
		2,800
West Virginia	15,000	
Wyoming	20,000	4,000

Note: Chapter 7 bankruptcy exemption levels are reported by state from Elias, Renauer, and Leonard (1994). The homestead exemption is for equity in owner-occupied principal residences. The personal property exemption is the sum of the state's exemptions for cash, equity in vehicles, and near-cash assets such as jewelry.

median value of the personal property exemption is \$7,000, with a minimum value of \$2,800. The maximum personal property exemption is \$60,000 in Texas, but only two other small states besides Texas have personal property exemptions higher than \$25,000. The correlation between the two exemptions is .24 but falls to .094 if Texas is excluded. 12

We enter the homestead exemption as dollar value (in thousands of dollars) and dollar value squared. For states with unlimited homestead exemptions, we set the homestead exemption equal to the maximum dollar value across all states, which is \$160,000. We also enter a separate dummy variable that equals one for states that have unlimited homestead exemptions. This means that the coefficient of the unlimited exemption dummy captures the marginal effect of the homestead exemption being unlimited rather than \$160,000. We enter the personal property exemption as dollar value (in thousands of dollars) and dollar value squared. Since no states have unlimited personal property exemptions, there is no unlimited dummy.

An important issue is whether the bankruptcy exemption level can be treated as exogenous to lenders' decisions to ration credit. As part of the 1978 Bankruptcy Code, Congress adopted a uniform federal bankruptcy exemption, but it gave the states the right to opt out of the federal exemption by adopting their own exemptions. All states did so by 1983, although about onethird allowed their residents to choose between the state's exemption and the federal exemption.¹³ Between 1983 and 1993, only a few states changed their exemption levels each year, and the federal bankruptcy exemptions remained unchanged. Exemption levels do not appear to be correlated with state loan market or demographic characteristics: a recent study explaining state exemption levels in the 1990s found that the only significant correlate was the state's exemption level in the 1920s (Posner, Hynes, and Malani, 2001). ¹⁴ We therefore treat exemption levels as exogenous.

Other important variables are measures of the owner's and the firm's previous financial difficulties. A previous bankruptcy filing has a theoretically ambiguous effect on creditworthiness. Because debtors who have filed for bankruptcy under Chapter 7 cannot file again under Chapter 7 for six years, they are more creditworthy from a lender's viewpoint. But debtors can still file for bankruptcy under Chapter 13 or default on repaying their loans without filing for bankruptcy. If they do the latter, then lenders have the right to sue for repayment, but such lawsuits are often not financially worthwhile. In addition, a past bankruptcy filing is evidence of weak entrepreneurial ability. These considerations suggest that lenders may view a past bankruptcy filing as negative evidence of creditworthiness. In contrast, a past personal or business delinquency is unambiguously negative evidence concerning creditworthiness. It signals weak entrepreneurial ability and also signals that debtors are knowledgeable about default procedures and willing to use them. Our bankruptcy variable is a dummy variable that equals one if the firm or its principal owner filed for bankruptcy within the past seven years. Our delinquency variables are separate dummy variables for whether owners have been delinquent on personal financial obligations during the past three years and whether the firm has been delinquent on business obligations during the past three years.¹⁵

¹² In computing a dollar value for the personal property exemption, we ignore exemptions for specific goods and use only exemptions for cash, vehicles, and near-cash assets such as jewelry that are specified as dollar values. Two states have unlimited exemptions for a single vehicle: Louisiana (for a "nonluxury auto") and Hawaii (for an otherwise unspecified vehicle). We code these at \$20,000 and \$40,000, respectively.

¹³ Sixteen states allowed their residents to choose between the state and the federal exemptions as of 1993. We substitute the federal exemption for the state exemption if the firm is located in a state that allows its residents to choose and the federal exemption is higher. Some states also have higher exemptions for married couples who file for bankruptcy, and, if so, we use the exemptions applicable to married couples.

¹⁴ The authors also ran a regression explaining state exemption levels in the 1920s and found that the population density and the proportion of farmers at that time were statistically significant. They interpret this result as indicating that debt relief was a more salient political issue in the past than it is now.

¹⁵ The personal financial delinquency variable equals one if the owner has been delinquent by 60 days or more on one or more personal obligations. The business delinquency variable is similarly defined. All of the bankruptcy and financial delinquency variables in our data are self-reported. But lenders can and routinely do obtain information on past bankruptcy filings and past financial delinquencies from credit reporting services. See White (1998a) for a model of the relationship between default and bankruptcy.

We include a number of firm demographic characteristics as additional evidence concerning firms' perceived creditworthiness. These are the firm's age, the owner's age, separate dummy variables for whether more than 50% of the firm's equity is owned by an African-American or a member of another minority group (Hispanic or Asian), separate dummy variables for whether the firm is family owned or female owned, and the firm's total employment in log form. Financial variables include the firm's ratio of debts to assets, its ratio of profits to assets, ¹⁶ the rate of growth of sales between 1990 and 1992, and a vector of dummy variables for the firm's sector (results for the latter are not shown). ¹⁷ We also include a dummy variable for a high Herfindahl index of bank deposit concentration in the market where the firm is located and the number of lenders that the firm borrows from. If the lender is a bank, we include the number of years that the firm and the bank have had a relationship, and we also include a dummy variable that equals one if the years-of-relationship variable is missing. We also include dummy variables for whether the firm has a checking or saving account with the bank and for whether the firm purchases services from the bank that are information intensive. ¹⁸ Finally, as a measure of local macroeconomic conditions, we include the unemployment rate in the state where the firm is located.

Table 2 shows summary statistics.¹⁹ Statistics are shown by type of firm and separately by whether firms are credit rationed (discouraged/denied). The overall probability of firms being credit rationed is 526/(526 + 1303) = .29 for noncorporate firms versus 736/(736 + 2072) = .26 for corporations. Firms that are credit rationed are much more likely to have owners who have filed for bankruptcy or have been financially delinquent in the past. The proportion of credit-rationed firms/owners that have filed for bankruptcy is .06-.08, compared to only .01 for non-credit-rationed firms/owners. The proportion of credit-rationed firms that have past business delinquencies is .36 for noncorporate firms and .45 for corporate firms, compared to only .11 and .13 for noncorporate and corporate firms that are not credit rationed, respectively. The figures for past personal delinquency have a similar pattern. Among the important differences between the two types of firms are that corporate firms are larger on average, less likely to be family owned, and less likely to be owned by African-Americans or other minorities.

□ **Credit rationing.** Table 3 shows the results of logit regressions explaining whether firms are credit rationed. Noncorporate firms are in the left-hand column and corporate firms are in the right-hand column.²⁰ Variables that are significant at the 95% level are marked with an asterisk. The dollar homestead exemption variable and the unlimited homestead exemption dummy have the predicted positive signs in both regressions, while the squared dollar homestead exemption variables are negative in both regressions.²¹

The personal property exemption coefficients are not statistically significant for either type of firm. This probably reflects the fact that the personal property exemption is smaller and less variable than the homestead exemption. In addition, the state that has by far the largest personal

¹⁶ Because the latest loan application can be any time during the past three years, financial variables such as the level of debt and the level of assets may or may not include the proceeds of the latest loan. See below for discussion of robustness checks in which we compare the results for firms that applied for loans more recently versus further back in time.

¹⁷ See Gordon and MacKie-Mason (1994) for discussion of firms' tax incentives to choose corporate versus noncorporate form. They show that owners of unprofitable firms have an incentive to choose noncorpoorate status, while owners of profitable firms have an incentive to choose the corporate form. We caputre the tax effrect on corporate form by including profit relative to assets as an explanatory variable.

 $^{^{18}}$ The "years of relationship missing" variable is included because otherwise we would have to drop 120 noncorporate and 50 corporate observations. We follow Petersen and Rajan (1994) in our choice of financial variables. Following their lead, we truncate sales growth at the 95% level and define the cutoff for a high Herfindahl index at HHI \geq 1800.

 $^{^{19}}$ In all calculations, we use the NSSBF sampling weights that make the sample representative of the target population of U.S. small businesses.

²⁰ Robust standard errors clustered by state are reported. See Moulton (1986).

²¹ The linear exemption term for corporations and the unlimited exemption dummy for noncorporates were not significant. However, an *F*-test shows that the three homestead exemption variables are jointly statistically significant at the .99 level in both samples.

TABLE 2 **Summary Statistics**

	Noncorporate				Corporate			
	Not Credit Constrained		Credit Constrained		Not Credit Constrained		Credit Constrained	
Variable	Mean	Standard Deviation	Mean	Standard Deviation	Mean	Standard Deviation	Mean	Standard Deviation
Discouraged/denied	.00	.00	1.00	.00	.00	.00	1.00	.00
Discouraged	.00	.00	.95	.22	.00	.00	.93	.25
Denied	.00	.00	.57	.50	.00	.00	.47	.50
Homestead exemption	61.86	57.18	66.73	56.69	56.14	57.68	55.74	56.16
Personal property exemption	12.64	16.27	12.84	17.00	11.61	14.47	10.83	13.73
Past bankruptcy filing	.01	.11	.06	.25	.01	.11	.08	.27
Past personal delinquency	.09	.28	.37	.48	.05	.22	.25	.43
Past business delinquency	.11	.31	.36	.48	.13	.33	.45	.50
Owner's age	50.65	11.83	46.51	10.18	51.57	11.27	47.99	10.63
Firm's age	15.43	13.77	10.95	8.57	17.26	14.61	12.76	12.19
Family owned	.89	.31	.94	.23	.75	.44	.77	.42
Female owned	.23	.42	.26	.44	.21	.41	.27	.44
African-American owned	.06	.24	.26	.44	.04	.20	.18	.39
Other minority owned	.23	.42	.51	.50	.13	.34	.32	.47
Employment	9.31	30.19	6.23	21.35	52.69	76.86	29.52	54.90
Profits/assets ratio	1.26	6.31	1.22	4.24	.49	2.81	.47	3.70
Debts/assets ratio	.53	.82	.70	.80	.60	.62	1.03	2.94
Sales growth	15.39	416.72	3.12	43.69	2.38	39.29	1.25	2.40
HHI > 1800	.53	.50	.47	.50	.49	.50	.42	.49
Years of bank relationship	9.83	8.82	6.91	5.95	10.00	9.39	6.72	6.52
Checking account at bank	.94	.25	.89	.32	.94	.24	.87	.34
Number of lenders	.03	.24	.06	.39	.19	.67	.27	.98
State unemployment rate	.07	.02	.07	.02	.07	.02	.07	.02
Loan size	759.92	5,539.52	308.73	1,234.63	1,406.30	5265.58	849.02	2,893.89
Interest rate	8.91	2.40	10.17	3.76	8.05	1.76	8.96	2.56
Number of observations	1,303		526		2,072		736	

Note: This table gives means and standard deviations of select variables from the 1993 National Survey of Small Business Finance. Noncorporate firms include proprietorships and partnerships; corporate firms include corporations and S-corporations. "Credit constrained" indicates that the firm was either discouraged from applying for a loan or applied and was denied, both within the past three years. All dollar amounts are in thousands

property exemption—Texas—also has an unlimited homestead exemption, so that the homestead exemption may be capturing part of the effect of the personal property exemption on lending behavior.²²

The past bankruptcy filing and past personal and business delinquency variables are all positive and significant for both types of firms. The sizes of the coefficients are similar in the two samples, suggesting that lenders view the effects of past bankruptcy and delinquency similarly for noncorporate and corporate borrowers. Although the theoretical prediction for the sign of the past bankruptcy variable is ambiguous, the results suggest that lenders view a past bankruptcy filing as a clear negative signal concerning creditworthiness. The positive signs of the past financial delinquency variables are also in line with our predictions. Important results for other variables, such as the fact that firms are more likely to obtain loans when they have had longer relationships

²² Studies of the effect of bankruptcy exemptions on whether loan applicants are denied mortgage loans have found positive and significant effects for the personal property exemption. But these studies use much larger datasets. See Berkowitz and Hynes (1999) and Lin and White (2001).

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with their lenders and the fact that African-American and minority-owned firms are more likely to be credit rationed, have been discussed in earlier articles.²³

We calculated the predicted probability of firms being credit rationed for various exemption levels. We assume that both types of firms are family owned, are not African-American or other minority owned, and have average values for the other right-hand-side variables. If a noncorporate firm is located in a state whose homestead and personal property exemptions are both at the 25th percentile of the relevant distributions, and if neither the firm nor its owner has previously filed for bankruptcy or been delinquent, then its probability of being credit rationed is .122. This figure rises to .140 if the homestead exemption in the firm's state is instead at the 50th percentile, and to .196 if the homestead exemption is at the 75th percentile.

Overall, the probability of noncorporate firms being credit rationed rises by 32% if firms are located in states with unlimited rather than low homestead exemptions. For corporate firms, the pattern is similar. The predicted probability of being credit rationed is .196 if both exemptions are at the 25th percentile, .221 if the homestead exemption rises to the 70th percentile, .172 if the exemption rises to the 80th percentile, and .255 if the exemption is unlimited. Overall, the probability of corporate firms being credit rationed rises by 30% if firms are located in states with unlimited rather than low homestead exemptions. These results are generally supportive of the theoretical model discussed in the previous section. Risk-averse entrepreneurs gain from having partial wealth insurance, and this causes demand for loans to rise even though interest rates are rising.

Nevertheless, the probability of credit rationing is not monotonically increasing in the homestead exemption level. For noncorporates, for example, the probability drops to .154 at the 90th percentile and then rises to .161 if the homestead exemption is unlimited. We do not have a good explanation for why the probability of credit rationing displays this nonmonotonic region when the homestead exemption level is not unlimited.

These results have several implications. First, corporations are more likely to be credit rationed than noncorporate firms over most of the observed range (though not when the homestead exemption is around the 90th percentile). This is in accord with our predictions, since owners' assets always guarantee the debt of noncorporate firms but do not always guarantee the debt of corporate firms. All else equal, more assets are available to back up loans to noncorporations. Second, the marginal effect of bankruptcy exemptions are larger for noncorporate than corporate firms. This makes sense because the sample of corporate firms is a mixture of corporations that are able to borrow on their own and corporations that are unable to borrow unless their owners provide guarantees. For the former group, the exemption level is irrelevant, and for the latter, it should have the same marginal effect as for the sample of noncorporations. Because the marginal effects of the exemption variables for corporations are an average over the two types, we expect them to be smaller than the marginal effects for noncorporate firms.

Holding the exemption levels fixed, the probability of firms being credit rationed approximately triples when their owners have previously filed for bankruptcy and approximately doubles when firms or their owners have previous financial delinquencies, with a previous business delinquency having a larger effect than a previous personal delinquency. These effects are similar for both noncorporate and corporate firms. Both types of firms have about a 50% probability of being credit rationed if their owners have previously filed for bankruptcy. The results suggest that a past bankruptcy filing or a past delinquency severely handicaps entrepreneurs who attempt to obtain loans for a new business.

We ran several additional checks to determine whether the results are robust to changes in the specification of the model. First, we checked whether the two types of credit-rationed firms—"discouraged" versus "denied"—differ systematically from each other. To do so, we reran the

²³ See Petersen and Rajan (1994) and Berger and Udell (1995) for discussion of the effect of borrower-lender relatinships. See Cole (1998), Cole, Goldberg, and White (2000), and Cavalluzzo, Cavalluzzo, and Wolken (2002) for discussion of the results for minorities and African-Americans. All of these articles used the same or earlier versions of the NSSBF.

TABLE 3 Logit Regressions Explaining Whether Firms Are Credit Rationed

	Noncorporate Firms		Corporate Firms	
	Coefficient	Standard Error	Coefficient	Standard
	Coefficient	EHOI	Coefficient	Error
Homestead exemption	.0186*	(.0074)	.0115	(.00625)
Homestead exemption ²	00011*	(.000045)	000120*	(.0000438)
Unlimited homestead exemption	.512	(.377)	1.69*	(.436)
Personal property exemption	.0118	(.0209)	.0104	(.0223)
Personal property exemption ²	000218	(.00033)	000281	(.000376)
Past bankruptcy	1.58*	(.444)	1.58*	(.364)
Past personal delinquency	.902*	(.219)	.791*	(.219)
Past business delinquency	1.24*	(.175)	1.16*	(.169)
Owner's age	0143*	(.0069)	00823	(.00568)
Firm's age	0103	(.00857)	00728	(.00664)
Family owned	.492*	(.254)	.0635	(.100)
Female owned	156	(.126)	.126	(.154)
African-American owned	1.20*	(.180)	1.008*	(.214)
Other minority owned	.548*	(.173)	.212	(.176)
Employment (in logs)	0750	(.0865)	176*	(.0604)
Profit/asset ratio	.0119	(.00857)	0493*	(.0210)
Debt/asset ratio	.197	(.110)	.265*	(.095)
Sales growth	000157	(.000227)	000453	(.00102)
HHI > 1800	0633	(.184)	0312	(.198)
Years of bank relationship	0264*	(.0142)	0240*	(.0117)
Years of bank relationship missing	195	(.310)	.297	(.584)
Checking account at bank	369	(.273)	700*	(.215)
Number of lenders	.356	(.215)	.0619	(.0612)
State unemployment rate	1.92	(3.93)	-1.415	(5.02)
Number of observations	1,8	01	2,7	780
Pseudo-R ²	.18	39	.1	60

Note: The dependent variable is whether firms were credit constrained (discouraged/denied) within the past three years. Asterisks indicate significance at the 95% level.

model in Table 3 for each group separately. Second, we redefined our benchmark credit-rationing variable so that firms were treated as credit rationed if they were denied credit on a prior loan application, even though their most recent application was approved, or if they were offered credit but on worse terms than they applied for. Third, firms' most recent loan application may occur anytime during the three years prior to the survey. Early loan applications present potential endogeneity problems, because the bankruptcy exemption variables and the other control variables are for the year of the survey. We reran the models in Table 3 but excluded firms whose most recent loan applications occurred more than one year prior to the survey. In all cases, the results for the exemption and financial delinquency variables remained substantially the same as those in Table 3.

The effect of bankruptcy exemptions on the probability of credit rationing should depend on the wealth of the potential borrower, since debtors are less likely to file for bankruptcy at a given exemption level when they have higher wealth. Because our dataset consists of small businesses, we do not know the business owner's personal wealth. But we do know the net value of business assets minus liabilities, and we use this as a proxy for entrepreneurs' wealth. We reran the model in Table 3 with interactions between all of the exemption variables and a dummy variable that equals one if business net assets are in the top half of the distribution.

The results are shown in Table 4 for the exemption variables only. Firms with relatively few net assets have exemption coefficients that are larger in absolute value and more statistically © RAND 2004.

Pseudo-R2

	Noncorporate Firms		Corporate Firms	
	Coefficient	Standard Error	Coefficient	Standard Error
Homestead exemption	.0248*	(.00793)	.0167*	(.00631)
Homestead exemption ²	000167*	(.000050)	000190*	(.000045)
Unlimited homestead exemption	.949*	(.413)	2.37*	(.626)
Homestead exemption × high asset dummy	0129*	(.00443)	00964	(.00651)
Homestead exemption $^2 \times$ high asset dummy	.000105*	(.000034)	.000127*	(.000057)
Unlimited homestead exemption × high asset dummy	947	(.567)	-1.25	(.850)

TABLE 4 Logit Regressions Explaining Whether Firms Are Credit Rationed When Net Assets Are High Versus Low

Note: The model is the same as in Table 3, except that the exemption variables are interacted with a dummy variable equaling one if the firm's net assets are in the top half of the distribution. Assets are net of liabilities. Only the results for the homestead exemption variables are given. Asterisks indicate significance at the 95% level.

significant than the results for the entire sample in Table 3. For firms with high net assets, the overall effect is close to zero—calculated by adding the coefficient for the low-asset firms plus the coefficient on the interaction terms. Thus, the exemption is more important as a determinant of credit rationing for firms with low net asset levels than for firms with high net asset levels.

This finding is as expected. Demand for loans is increasing in the exemption, regardless of whether the firm has high or low wealth. But while lenders accommodate the increased loan demand for customers whose businesses have high net assets; they respond by rationing credit to customers whose businesses have low net assets.²⁴

Interest rates. Table 5 gives the results of Tobit regressions explaining the interest rate on the most recent loan that the firm obtained.²⁵ The samples are the same as for the model of credit rationing in Table 3. All firms that did not receive credit are right censored at an interest rate of 17% (the maximum interest rate in the sample is 16.5%).²⁶ All of the homestead exemption variables are statistically significant in both samples, except for the unlimited homestead exemption dummy in the corporate sample. The results are similar to the credit-rationing results: the homestead exemption coefficients are positive, but the squared exemption coefficients are negative. Again, none of the personal property exemption variables are statistically significant. The past bankruptcy variables enter significantly positive as expected in both samples, but among the past delinquency variables, only the personal delinquency variable in the corporate sample is statistically significant. Among the other variables, an interesting result is that the high Herfindahl index variable and the number of lenders variable are significantly negative (though only at the 10% level for noncorporate firms). Small businesses pay lower interest rates when there is more competition in the local lending markets and when firms borrow from more lenders. Larger firms also pay lower interest rates, and firms owned by African-Americans and other minorities pay higher interest rates.

²⁴ See Gropp, Scholz, and White (1997) for discussion of the effect of borrowers' wealth on access to credit by households. They find similar differences in the probability of credit rationing for low-wealth versus high-wealth borrowers.

²⁵ We did not use a two-stage Heckman model because it did not seem possible to separate the credit-rationing model from the interest rate determination model, since any observable information that affects what interest rate lenders charge also affects whether they ration credit.

²⁶ Because we have no information on the interest rate for firms that were credit rationed, we reran the model with the truncated interest rate set at levels between 18% and 30%. The results for higher interest rates were that the size and significance of the homestead exemption coefficients gradually increased, but everything remained the same. We report the results for the interest rate of 17% because this specification has the highest log-likelihood value.

TABLE 5 Effects of Bankruptcy Exemptions on Interest Rates

	Noncorporate Firms		Corporate Firms	
	Coefficient	p-value	Coefficient	p-value
Homestead exemption	.0824*	.009	.0367*	.033
Homestead exemption ²	000588*	.007	000271*	.030
Unlimited homestead exemption	4.91*	.025	1.70	.168
Personal property exemption	139	.233	.0652	.283
Personal property exemption ²	.0011	.565	000939	.362
Past bankruptcy	5.71*	.052	2.20*	.053
Past personal delinquency	2.03	.077	1.45*	.021
Past business delinquency	867	.392	-0.540	.250
Owner's age	.106*	.002	.0887*	.000
Firm's age	0257	.467	.0191	.317
Family owned	2.28	.058	905*	.030
Female owned	.501	.541	.513	.189
African-American owned	5.26*	.027	8.31*	.000
Other minority owned	9.10*	.000	1.88*	.007
Employment (in log form)	-1.35*	.001	-1.33*	.000
Profit/asset ratio	.111	.171	.139*	.017
Debt/asset ratio	507	.177	206	.205
Sales growth	00104	.102	.00952	.442
HHI > 1800	-1.18	.088	854*	.015
Years of bank relationship	.000923	.986	00365	.989
Years of bank relationship missing	480	.769	3.95*	.007
Checking account at bank lender	1.31	.328	.844	.196
Number of lenders	-2.43	.058	939*	.000
Intercept	15.2*	.000	13.1*	.000
Number of observations		1,801		2,779
Pseudo-R ²		.0508		.0205
Log-likelihood	-2	2140.96	=:	5808.21

Note: This table presents results from Tobit regressions explaining the interest rate on loans for noncorporate and corporate firms. Interest rates for firms that were credit rationed are right censored 17%, which is above the maximum value of 16.5% in the sample. Asterisks indicate statistical significance at the 95% level.

We calculated the predicted effects of changes in the homestead exemption level on interest rates, following the same procedure as used above. If the homestead exemption increases from the 25th to 50th percentile, the interest rate for noncorporate firms is predicted to rise by .60 percentage points and, if the exemption increase is from the 50th to the 75th percentile, a further 1.33 percentage points. The overall increase in going from the 25th percentile to unlimited is a predicted increase of 2.15 percentage points. For corporate firms, the interest rate rises by .27 percentage points when the exemption level increases from the 25th to the 50th percentile, and by a further .56 percentage points from the 50th to the 75th percentile. For both firm types, however, the increase in interest rates is nonmonotonic when exemptions are around their 80th–90th percentiles. This result is surprising but consistent with the credit-rationing results.

As expected, the responsiveness of the interest rate to changes in the exemption level is higher for noncorporate firms than corporations. This is presumably because lenders are willing to lend to some small corporate firms qua corporations, but treat other small corporate firms as though they were noncorporate—i.e., they require that owners personally guarantee the loans and therefore bankruptcy exemption levels matter. A past bankruptcy filing also has a larger effect on noncorporate than corporate firms: it is predicted to increase the interest rate for noncorporate firms by 5.4 percentage points and for corporate firms by 2.1 percentage points.

TABLE 6 Effects of Bankruptcy Exemptions on Loan Size

	Noncorporate Firms		Corporate Firms	
	Coefficient	p-value	Coefficient	p-value
Homestead exemption	0970*	.010	0382	.077
Homestead exemption ²	.000677*	.010	.000304	.051
Unlimited homestead exemption	-5.18*	.050	-2.34	129
Personal property exemption	.126	.371	0802	.290
Personal property exemption ²	00082	.723	.00114	.376
Past bankruptcy	-7.56*	.034	-2.99*	.036
Past personal delinquency	-3.07*	.027	-1.93*	.014
Past business delinquency	1.08	.371	.730	.213
Owner's age	127*	.002	120*	.000
Firm's age	.0333	.434	0153	.519
Family owned	-2.74	.057	1.27*	.015
Female owned	-1.13	.247	782	.110
African-American owned	-5.26	.061	-10.04*	.000
Other minority owned	-10.7^{*}	.000	-2.35^{*}	.007
Employment (in log form)	1.68 *	.000	1.95*	.000
Profit/asset ratio	164	.109	173*	.017
Debt/asset ratio	1.02*	.016	.284	.158
Sales growth	.000927	.209	00851	.576
HHI > 1800	1.74*	.036	1.047*	.017
Years of bank relationship	0209	.737	00436	.898
Years of bank relationship missing	-9.99	.000	-5.89	.002
Checking account at bank lender	-1.363	.400	-1.34	.100
Number of lenders	2.83	.068	1.17*	.000
Intercept	2.28	.495	4.98*	.002
Number of observations	1,801		2,780	
Pseudo-R ²		.0485	.0220	
Log-likelihood				5133.26

Note: This table presents results from Tobit regressions explaining loan size (in log form) for noncorporate and corporate firms. Loan size for all firms that did not receive credit is censored at zero. p-values are given, and asterisks indicate statistical significance at the 95% level.

Loan size. For the most recent loan application, the NSSBF asks the size of the loan that the lender offered.²⁷ We ran Tobit models explaining loan size (in logs) for the samples of noncorporate and corporate firms where loan size is left censored at zero. Again, we expect the coefficients to reflect a combination of supply and demand considerations. An increase in the bankruptcy exemption level causes entrepreneurs to demand larger loans because the consequences of business failure are not as bad, but it also causes lenders to reduce loan supply because lending is more risky. Thus we could observe either positive or negative coefficients on the exemption variables, depending on which effect is larger.

The results are shown in Table 6. For both types of firms, the signs of the homestead exemption variables mirror those from the previous results—loan size is decreasing in homestead exemption. All three variables are statistically significant for noncorporate firms, but for corporations the unlimited dummy is not. The personal property exemption variables are not significant in either sample. As expected, the past bankruptcy filing and past personal delinquency variables are

²⁷ The NSSBF actually asks separately how much the firm applied for and how much the lender offered. In theory, these separate measures might allow us to separately estimate a demand curve from the former and a supply curve from the latter. In practice, however, the two variables are extremely closely related, with a correlation coefficient of .994. Presumably, firms apply for the amount of credit that they expect lenders to provide, and lenders may tell borrowers in advance how much they are willing to lend.

negative and significant in both regressions, but the past business delinquency variables are insignificant.

Finally, we turn to the predicted effects of the homestead exemption level on loan size. For noncorporate firms, if the exemption increases from the 25th to the 75th percentile, loan size falls by about \$79,000. The overall change predicted by an increase of the exemption from the 25th percentile to unlimited is a reduction of \$198,000. For corporate firms, if the exemption increases from the 25th to the 75th percentile, loan size is predicted to fall by \$68,000 and, when increased from the 25th percentile to unlimited, by \$41,000. The generally decreasing relationship between loan size and the exemption level is consistent with our model, though the nonmonotonic region is not. Finally, as predicted, the marginal effects of changes in the exemption level are larger for noncorporate firms than for corporations. A past personal bankruptcy filing is predicted to reduce loan size for noncorporate firms by \$470,000 and for corporations by \$72,000.

6. Conclusion

This article investigates how personal bankruptcy law affects small firms' access to credit. We show that higher personal bankruptcy exemptions are predicted to cause increased credit rationing and higher interest rates. These predictions apply to both noncorporate firms and corporations, although the marginal effects should be larger for noncorporate firms. Personal bankruptcy exemptions affect credit markets for small corporate firms because small business owners can easily transfer funds from the firm to themselves even if the firm is corporate. Thus lenders may not view the corporate/noncorporate distinction as meaningful for small firms.

We test the model and find that small businesses located in states with high homestead exemptions are more likely to be credit rationed, and if they receive loans, interest rates are higher and loan size is smaller. However, we find that the effect of increases in exemption levels is nonmonotonic: as the exemption level rises, the probability of credit rationing and the interest rate first rise, then fall, and then rise again when the exemption level becomes unlimited. For loan size, the pattern is the same with the opposite sign.

We also find that the effect of changes in exemption levels on the probability of credit rationing is much larger for firms with low net assets than for firms with high net assets, regardless of whether they are corporate or noncorporate. Notably, when owners of either type of firm have previously filed for bankruptcy, the probability that their firms are credit rationed triples, and when firms or their owners have previously been delinquent on personal or business obligations, the probability that firms are credit rationed doubles.

Each year for the past several years, one or both houses of the U.S. Congress have adopted legislation that would limit homestead exemptions to a maximum of \$100,000 or \$125,000. Although the intended effect of the change is to discourage well-off consumers from taking advantage of bankruptcy to shield their assets from creditors, our results suggest that the change would make it easier for small businesses to obtain credit.

References

BERGER, A.N. AND UDELL, G.F. "Relationship Lending and Lines of Credit in Small Firm Finance." Journal of Business, Vol. 68 (1995), pp. 351-382.

BERKOWITZ, J. AND HYNES, R. "Bankruptcy Exemptions and the Market for Mortgage Loans." Journal of Law and Economics, Vol. 42 (1999), pp. 809-830.

BESTER, H. "The Role of Collateral in a Model of Debt Renegotiation." Journal of Money, Credit and Banking, Vol. 26 (1994), pp. 72-86.

CAVALLUZZO, K.S., CAVALLUZZO, L.C., AND WOLKEN, J.D. "Competition, Small Business Financing, and Discrimination: Evidence from a New Survey." Journal of Business, Vol 75 (2002), pp. 641-679.

Cole, R.A. "The Importance of Relationships to the Availability of Credit." Journal of Banking and Finance, Vol. 22 (1998), pp. 959–977.

AND WOLKEN, J.D. "Financial Services Used by Small Businesses: Evidence from the 1993 National Survey of Small Business Finances." Federal Reserve Bulletin, Vol. 81 (1995), pp. 629–667.

© RAND 2004.

- , GOLDBERG, L.G., AND WHITE, L.J. "Cookie-Cutter Versus Character: The Microstructure of Small Business Lending by Large and Small Banks." In J.L. Blanton, A. Williams, and S.L.W. Rhine, eds., *Business Access to Capital and Credit*. Washington, D.C.: Federal Reserve, 2000.
- ELIAS, S., RENAUER, A., AND LEONARD, R. How to File for Bankruptcy, 4th ed. Berkeley, Calif.: Nolo Press, 1993.
- FAN, W. AND WHITE, M.J. "Personal Bankruptcy and the Level of Entrepreneurial Activity." *Journal of Law and Economics*, Vol. 46 (2003), pp. 543–568.
- GORDON, R.H. AND MACKIE-MASON, J.K. "Tax Distortions to the Choice of Organizational Form." *Journal of Public Economics*, Vol. 55 (1994), pp. 279–306.
- GROPP, R., SCHOLZ, J.K., AND WHITE, M.J. "Personal Bankruptcy and Credit Supply and Demand." *Quarterly Journal of Economics*, Vol. 112 (1997), pp. 217–252.
- HART, O. AND MOORE, J. "Default and Renegotiation: A Dynamic Model of Debt." *Quarterly Journal of Economics*, Vol. 113 (1998), pp. 1–41.
- LIN, E.Y. AND WHITE, M.J. "Bankruptcy and the Market for Mortgage and Home Improvement Loans." *Journal of Urban Economics*, Vol. 50 (2001), pp. 138–162.
- LONGHOFER, S.D. "Absolute Priority Rule Violations, Credit Rationing, and Efficiency." *Journal of Financial Intermediation*, Vol. 6 (1997), pp. 249–267.
- MOULTON, B.R. "Random Group Effects and the Precision of Regression Estimates." *Journal of Econometrics*, Vol. 32 (1986), pp. 385–397.
- Petersen, M.A. and Rajan, R.G. "The Benefits of Lending Relationships: Evidence from Small Business Data." *Journal of Finance*, Vol. 49 (1994), pp. 3–37.
- POSNER, E., HYNES, R., AND MALANI, A. "The Political Economy of Property Exemption Laws." Working paper, University of Chicago, 2001.
- Scott, J.A. AND SMITH, T.C. "The Effect of the Bankruptcy Reform Act of 1978 on Small Business Loan Pricing." *Journal of Financial Economics*, Vol. 16 (1986), pp. 119–140.
- STEINGOLD, F.S. Legal Guide for Starting and Running a Small Business. Berkeley, Calif.: Nolo Press, 1999.
- STIGLITZ, J.E. AND WEISS, A. "Credit Rationing in Markets with Imperfect Information." *American Economic Review*, Vol. 71 (1981), pp. 393–410.
- Sullivan, T.A., Warren, E., and Westbrook, J.L. As We Forgive Our Debtors: Bankruptcy and Consumer Debt in America. New York: Oxford University Press, 1989.
- U.S. SMALL BUSINESS ADMINISTRATION, OFFICE OF ADVOCACY. Small Business Growth by Major Industry, 1988–1995, 1998
- WHITE, M.J. "Why Don't More Households File for Bankruptcy?" *Journal of Law, Economics, and Organization*, Vol. 14 (1998a), pp. 205–231.
- ------. "Why It Pays to File for Bankruptcy: A Critical Look at Incentives Under U.S. Bankruptcy Laws and a Proposal for Change." *University of Chicago Law Review*, Vol. 65 (1998b), pp. 685–732.