

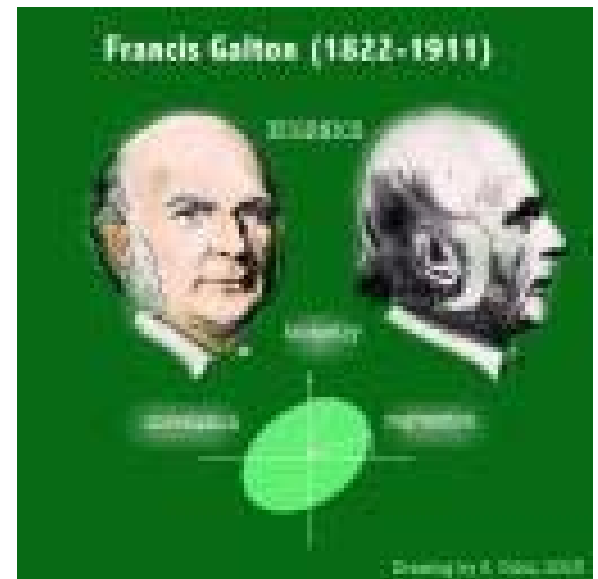
Who Invented Regression?

1. **Who invented regression?**
2. **Omitted Variables and Multivariate Regression**
3. **Omitted Variable Bias (OVB)**
4. **Experiments vs. OVB**
5. **R^2**



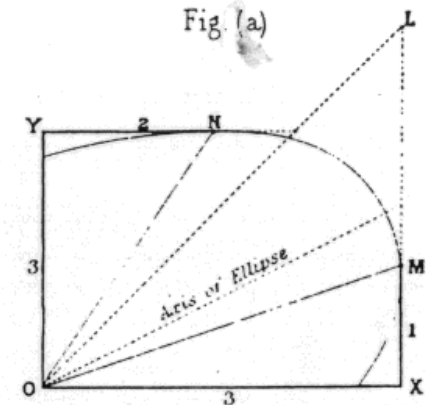
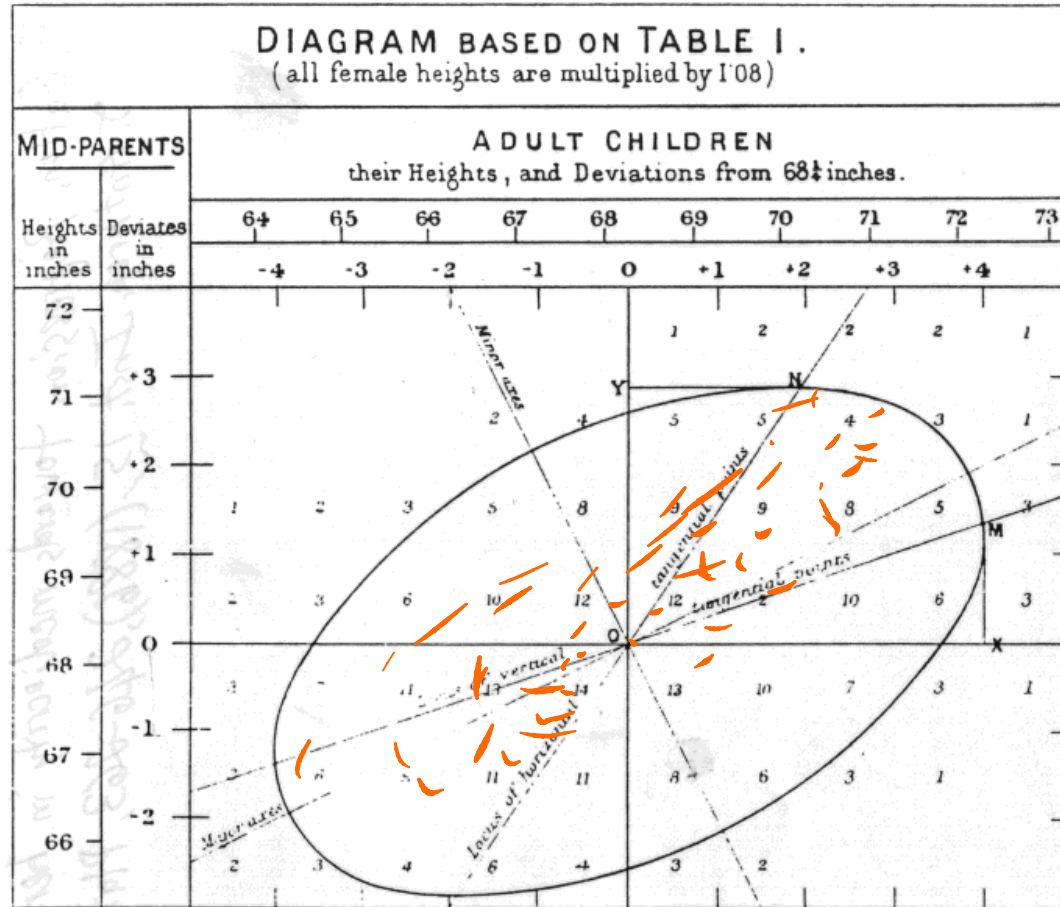
1. Who invented regression?

- Francis Galton,
 - climatologist,
 - gentleman explorer
 - social scientist



Heredity and Height

“regression” to the mean



Journal Anthropology Inst., Vol. XV, Pt. 1

2. Omitted Variables and Omitted Variable Bias

- **What if you left out an important variable?**
- **Many interesting relationships have more than 2 dimensions**

GRE prep course example

Coffee example

Problem set and exam example

- **We need more variables..**
“multivariate” regression

2. OLS Multivariate regression

The OLS Estimators, Predicted Values, and Residuals in the Multiple Regression Model

The OLS estimators $\hat{\beta}_0, \hat{\beta}_1, \dots, \hat{\beta}_k$ are the values of b_0, b_1, \dots, b_k that minimize the sum of squared prediction mistakes $\sum_{i=1}^n (Y_i - b_0 - b_1 X_{1i} - \dots - b_k X_{ki})^2$. The OLS predicted values \hat{Y}_i and residuals \hat{u}_i are:

$$\hat{Y}_i = \hat{\beta}_0 + \hat{\beta}_1 X_{1i} + \dots + \hat{\beta}_k X_{ki}, \quad i = 1, \dots, n, \text{ and} \quad (5.11)$$

$$\hat{u}_i = Y_i - \hat{Y}_i, \quad i = 1, \dots, n. \quad (5.12)$$

The OLS estimators $\hat{\beta}_0, \hat{\beta}_1, \dots, \hat{\beta}_k$ and residual \hat{u}_i are computed from a sample of n observations of $(X_{1i}, \dots, X_{ki}, Y_i)$, $i = 1, \dots, n$. These are estimators of the unknown true population coefficients $\beta_0, \beta_1, \dots, \beta_k$ and error term, u_i .

Look familiar? Same criterion with more variables.

2. Properties of OLS estimators in Multivariate Regression

- Consistent
- Unbiased
- Approximately $N(\cdot)$ in large samples
- Same first order conditions (for 2 or more X 's)

$$\sum_{i=1}^N e_i = 0$$

$$\sum_{i=1}^N X_{1i} e_i = 0$$

$$\sum_{i=1}^N X_{2i} e_i = 0$$

First order conditions for multivariate regression

$$\left. \begin{aligned} \frac{\partial e_i^2}{\partial b_0} &= \frac{\partial e_i^2}{\partial e_i} \times \frac{\partial e_i}{\partial b_0} \\ &= 2e_i (-1) \end{aligned} \right\}$$

$$\text{Min } \sum e_i^2, \quad e_i = (Y_i - \hat{Y}_i)$$

$$\{b_0, b_1, b_2, \dots, b_k\} \quad = (Y_i - \underline{b_0} - \underline{b_1}X_{1i} - \underline{b_2}X_{2i} - \underline{b_3}X_{3i} - \dots - \underline{b_k}X_{ki})$$

$k+1$ first order conditions..

$$0 = \frac{\partial \sum e_i^2}{\partial b_0} = \frac{\partial e_1^2}{\partial b_0} + \frac{\partial e_2^2}{\partial b_0} + \dots + \frac{\partial e_n^2}{\partial b_0} = -2e_1 - 2e_2 - 2e_3 - 2e_4 \dots - 2e_n$$

$$= -2 \sum e_i \Leftrightarrow \boxed{\sum e_i = 0} \quad \#1$$

$$0 = \frac{\partial \sum e_i^2}{\partial b_1} = \frac{\partial e_1^2}{\partial b_1} + \frac{\partial e_2^2}{\partial b_1} + \frac{\partial e_3^2}{\partial b_1} + \dots + \frac{\partial e_n^2}{\partial b_1} = -2e_1 X_{11} - 2e_2 X_{12} - 2e_3 X_{13} \dots - 2e_n X_{1n}$$

$$= -2 \sum e_i X_{1i} \Leftrightarrow \boxed{\sum e_i X_{1i} = 0} \quad \#2$$

$$\vdots$$

$$0 = \frac{\partial \sum e_i^2}{\partial b_k} = -2 \sum e_i X_{ki} \Leftrightarrow \boxed{\sum e_i X_{ki} = 0} \quad \#k+1$$

If the $k+1$ equations are not linearly dependent we can solve for $b_0, b_1, b_2, \dots, b_k$.

In STATA:
regress y x1 x2 .. xk, robust

3. Omitted Variable “Bias”

- Short regression

$$\rightarrow y = b_0^s + \underline{b_1^s} x_1 + e^s \quad (\text{SR})$$

eg. y - SAT score
 x_1 - SAT prep.
 x_2 - SAT ability

- Long regression

$$\rightarrow y = b_0^L + b_1^L x_1 + b_2^L x_2 + e^L \quad (\text{LR})$$

- Claim:

$$b_1^s = b_1^L + b_2^L b_{21},$$

$$\frac{dy}{dx_1} = \frac{\partial y}{\partial x_1} \Big|_{x_2} + \frac{\partial y}{\partial x_2} \frac{dx_2}{dx_1}$$

b_{21} is slope of a regression of x_2 on x_1

$$b_{21} = \frac{\sum (x_{1i} - \bar{x}_1)(x_{2i} - \bar{x}_2)}{\sum (x_{1i} - \bar{x}_1)^2}$$

Omitted variable bias formula

- derivation

$$\begin{aligned}
 b_1^S &= \frac{\sum (x_{1i} - \bar{x}_1)(y_i - \bar{y})}{\sum (x_{1i} - \bar{x}_1)^2} = \frac{\sum (x_{1i} - \bar{x}_1) y_i}{\sum (x_{1i} - \bar{x}_1)^2} = \frac{\sum (x_{1i} - \bar{x}_1) [b_0^L + b_1^L x_{1i} + b_2^L x_{2i} + e_i]}{\sum (x_{1i} - \bar{x}_1)^2} \\
 &= \overset{1 \downarrow}{0} + b_1^L \frac{\sum (x_{1i} - \bar{x}_1) x_{1i}}{\sum (x_{1i} - \bar{x}_1)(x_{1i} - \bar{x}_1)} + b_2^L \frac{\sum (x_{1i} - \bar{x}_1)(x_{2i} - \bar{x}_2)}{\sum (x_{1i} - \bar{x}_1)^2} +
 \end{aligned}$$

$$b_1^S = b_1^L + b_2^L b_{21}$$

$$\frac{\sum (x_{1i} - \bar{x}_1) e_i^L}{\sum (x_{1i} - \bar{x}_1)^2}$$

$$\left. \begin{aligned}
 &\sum (x_{1i} - \bar{x}_1) e_i^L \\
 &= \sum x_{1i} e_i^L - \sum \bar{x}_1 e_i^L = 0
 \end{aligned} \right|$$

4. Why experiments eliminate OVB

$$(s) Y = b_0^s + b_1^s X_1 + e^s$$

$$(L) Y = b_0^L + b_1^L X_1 + b_2^L X_2 + e^L$$

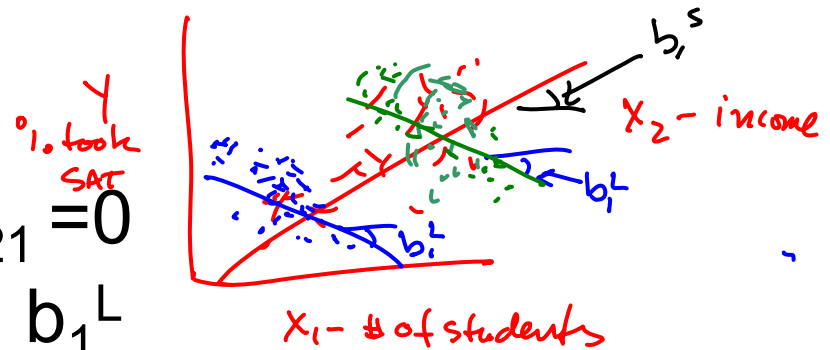
$$\begin{matrix} (+) & & (-) & & (+) & & (+) \\ b_1^s = & b_1^L + & b_2^L & b_{21}, \\ \underline{\quad} & & \underline{\quad} & \underline{\quad} \\ & & & 0 \end{matrix}$$

- So there's no OVB if $b_{21} = 0$
i.e., $b_{21} = 0$ implies $b_1^s = b_1^L$

.. Which you can guarantee if you design an experiment in which X_1 is uncorrelated with other X 's (omitted variables).

Random assignment of X_1 is sure to do that.

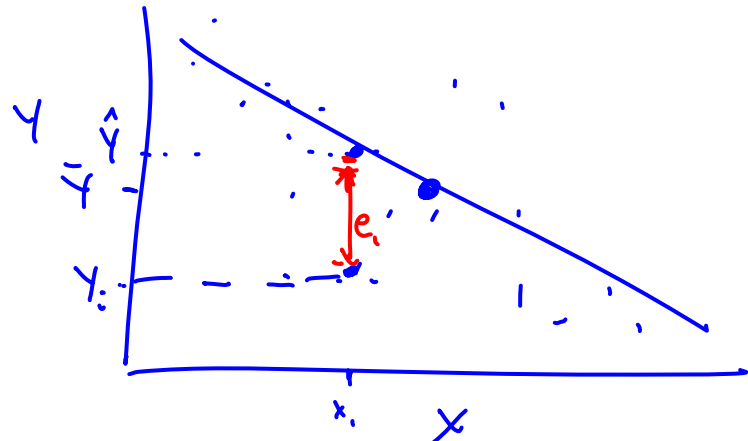
.. Back to examples to demonstrate



5. R^2 – How much Variation Explained?

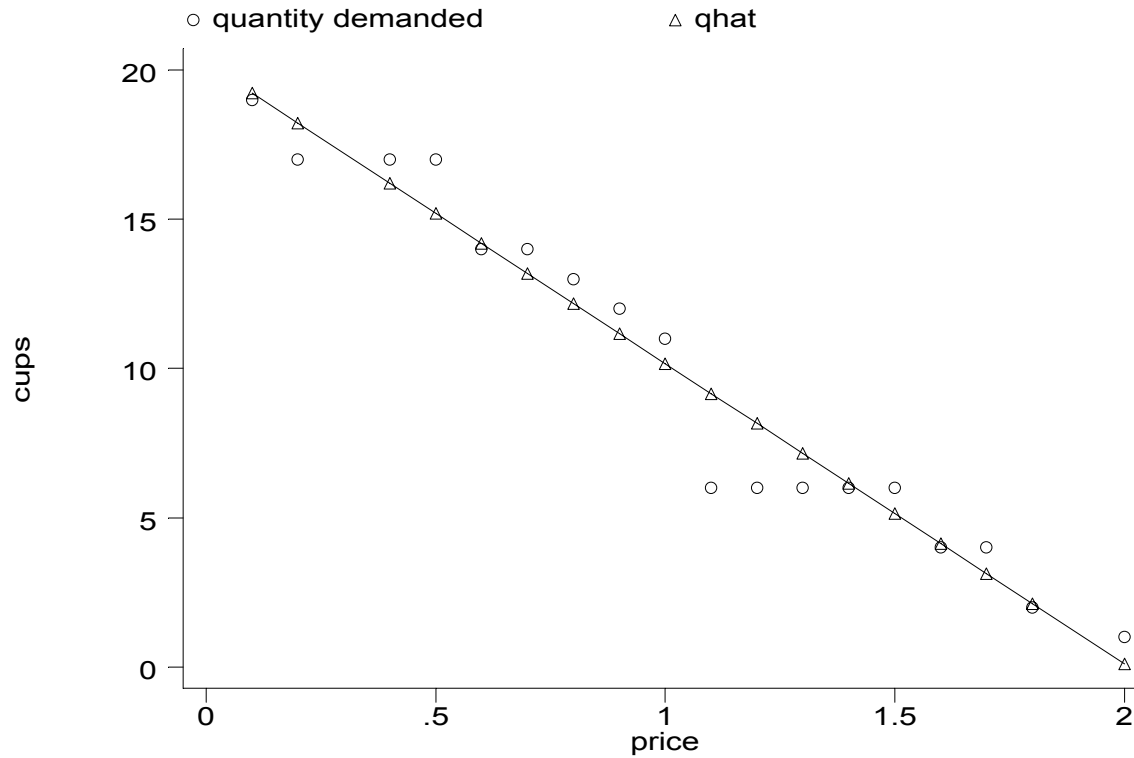
- How much of the variation in Y did we explain with the regression line?

$$R^2 = \frac{\sum_{i=1}^N (\hat{y}_i - \bar{y})^2}{\sum_{i=1}^N (y_i - \bar{y})^2}$$
$$= 1 - \frac{\sum_{i=1}^N e_i^2}{\sum_{i=1}^N (y_i - \bar{y})^2}$$



CLAIM: $\sum (y_i - \bar{y})^2 = \sum (\hat{y}_i - \bar{y})^2 + \sum e_i^2$
total = on the line + off the line

Coffee Demand – High R^2



E.g. Coffee Demand – high R^2

- p is the price of coffee,
- q is the quantity (in cups)

```
. reg q p, robust
```

```
Regression with robust standard errors
```

```
Number of obs =      23  
F( 1, 21) =      28.20  
Prob > F      =      0.0000  
R-squared     =      0.7349  
Root MSE     =      4.0549
```

```
-----  
              |  
              |      Robust  
              |      Coef.  Std. Err.      t    P>|t|     [95% Conf. Interval]  
-----+-----  
      q |      -6.246766   1.176301   -5.31   0.000   -8.693018   -3.800513  
      _cons |      17.5064   1.822797    9.60   0.000   13.71568   21.29711  
-----
```

Eg. Wage Regression - Low R²

* Lhwage is log(hourly wage), ed is years of education

```
regress lhwage ed, robust
```

Regression with robust standard errors

```
Number of obs = 13743  
F( 1, 13741) = 1795.40  
Prob > F      = 0.0000  
R-squared     = 0.1185  
Root MSE     = .5083
```

		Robust				
lhwage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
ed	.0704563	.0016628	42.37	0.000	.0671969	.0737156
_cons	.9852746	.0238393	41.33	0.000	.9385464	1.032003
